Law of Large Numbers for the SIR model

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- Here we assume that each infected individual infects at rate λ during a period of length $\mathcal{I}_{\rangle} \sim \textit{Exp}(\mu)$, where the \mathcal{I}_{\rangle} 's are i.i.d..
- The model for the evolution of $(S^N(t), I^N(t), R^N(t))$ is the following

$$S^{N}(t) = S^{N}(0) - \int_{0}^{t} \int_{0}^{\lambda \bar{S}^{N}(s^{-})I^{N}(s^{-})} Q_{inf}(ds, du),$$

$$I^{N}(t) = I^{N}(0) + \int_{0}^{t} \int_{0}^{\lambda \bar{S}^{N}(s^{-})I^{N}(s^{-})} Q_{inf}(ds, du) - \int_{0}^{t} \int_{0}^{\mu I^{N}(s^{-})} Q_{rec}(ds, du)$$

$$R^{N}(t) = R^{N}(0) + \int_{0}^{t} \int_{0}^{\mu I^{N}(s^{-})} Q_{rec}(ds, du).$$

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We divide the above system by N, yielding

$$ar{S}^N(t) = S^N(0) - \lambda \int_0^t ar{S}^N(s) ar{I}^N(s) ds - arepsilon_{inf}^N(t),$$
 $ar{I}^N(t) = I^N(0) + \lambda \int_0^t ar{S}^N(s) ar{I}^N(s) ds + arepsilon_{inf}^N(t) - \mu \int_0^t I^N(s) ds - arepsilon_{rec}^N(t),$ $ar{R}^N(t) = R^N(0) + \mu \int_0^t I^N(s) ds + arepsilon_{rec}^N(t),$

where

$$arepsilon_{inf}^{N}(t) = rac{1}{N} \int_{0}^{t} \int_{0}^{\lambda \bar{S}^{N}(s^{-})I^{N}(s^{-})} \bar{Q}_{inf}(ds, du)$$
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• Clearly $\mathbb{E}[|\varepsilon_{inf}^N(t)|^2] \leq \frac{\lambda^2}{N} \int_0^t \bar{S}^N(s) \bar{I}^N(s) ds \leq \frac{\lambda^2 t}{N}, \ \mathbb{E}[|\varepsilon_{rec}^N(t)|^2] \leq \frac{t}{N}.$

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- We now write the VI N-model. First we are given $(S^N(0), I^N(0), R^N(0))$, such that $S^N(0) + I^N(0) + R^N(0) = N$.
- To the initially infected individuals, we associate $\{\lambda_j^0(t),\ t\geq 0,\ 1\leq j\leq I^N(0)\}$, and define $\eta_j^0=\sup\{t,\lambda_j^0(t)>0\}$
- Initially susceptible individuals get infected at times τ_i^N , $1 \le i \le S^N(0)$. To each i, we associate $\lambda_i(t)$, which is the infectivity at time t after its infection time. Hence at time t, the infectivity of individual i is $\lambda_j(t-\tau_i^N)$. We assume that $\lambda_i(t)=0$ for t<0. We define $\eta_i=\sup\{t>0,\ \lambda_i(t)>0\}$.
- We assume that the two sequences $\{\lambda_j^0(\cdot),\ 1\leq j\leq I^N(0)\}$ and $\{\lambda_i(\cdot),\ 1\leq i\leq S^N(0)\}$ are both i.i.d., and mutually independent. Also $\lambda^0,\lambda\in D,\ 0\leq \lambda_1^0(t)\vee\lambda_1(t)\leq \lambda^*,\ \forall t\geq 0$ a.s. (where λ^* is a constant). We let $F^0(t)=\mathbb{P}(\eta_1^0\leq t),\ F(t)=\mathbb{P}(\eta_1\leq t),\ F_0^c(t)=1-F_0(t),\ F^c(t)=1-F(t).$

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- We first write $S^N(t) = S^N(0) A^N(t)$.
- The total force of infection at time t is

$$\mathfrak{F}^{N}(t) = \sum_{j=1}^{IN(0)} \lambda_{j}^{0}(t) + \sum_{i=1}^{A^{N}(t)} \lambda_{i}(t\tau_{i}^{N}),$$

$$A^{N}(t) = \int_{0}^{t} \int_{0}^{\bar{S}^{N}(s^{-})} \mathfrak{F}^{N}(s^{-})} Q(ds, du).$$

Moreover

$$I^{N}(t) = \sum_{j=1}^{I^{N}(0)} 1_{\eta_{j}^{0} > t} + \sum_{i=1}^{S^{N}(0)} 1_{\tau_{i}^{N} + \eta_{i} > t},$$
 $R^{N}(t) = R^{N}(0) + \sum_{j=1}^{I^{N}(0)} 1_{\eta_{j}^{0} \le t} + \sum_{i=1}^{S^{N}(0)} 1_{\tau_{i}^{N} + \eta_{i} \le t}.$

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$$\begin{split} I^N(t) &= \sum_{j=1}^{I^N(0)} 1_{\eta^0_j > t} + \sum_{i=1}^{S^N(0)} 1_{\tau^N_i + \eta_i > t}, \\ R^N(t) &= R^N(0) + \sum_{j=1}^{I^N(0)} 1_{\eta^0_j \le t} + \sum_{i=1}^{S^N(0)} 1_{\tau^N_i + \eta_i \le t}. \end{split}$$

Law of Large Numbers

• Notation : $\bar{X}^N(t) = X^N(t)/N$. As $N \to \infty$, in probability, loc. uniformly in t, $(\bar{S}^N(t), \bar{\mathfrak{F}}^N(t), \bar{I}^N(t), \bar{R}^N(t)) \to (\bar{S}(t), \bar{\mathfrak{F}}(t), \bar{I}(t), \bar{R}(t))$, where

$$\bar{S}(t) = \bar{S}(0) - \int_0^t \bar{S}(s)\bar{\mathfrak{F}}(s)ds,$$

$$\bar{\mathfrak{F}}(t) = \bar{I}(0)\bar{\lambda}^0(t) + \int_0^t \bar{\lambda}(t-s)\bar{S}(s)\bar{\mathfrak{F}}(s)ds.$$

and

$$\begin{split} \bar{I}(t) &= \bar{I}(0)F_0^c(t) + \int_0^t F^c(t-s)\bar{S}(s)\bar{\mathfrak{F}}(s)ds, \\ \bar{R}(t) &= \bar{R}(0) + \bar{I}(0)F_0(t) + \int_0^t F(t-s)\bar{S}(s)\bar{\mathfrak{F}}(s)ds \end{split}$$

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$$ar{I}(t) = ar{I}(0)F_0^c(t) + \int_0^t F^c(t-s)ar{S}(s)ar{\mathfrak{F}}(s)ds, \ ar{R}(t) = ar{R}(0) + ar{I}(0)F_0(t) + \int_0^t F(t-s)ar{S}(s)ar{\mathfrak{F}}(s)ds.$$

Another formulation of the stochastic model

We can rewrite

$$S^{N}(t) = S^{N}(0) - \sum_{i=1}^{S^{N}(0)} A_{i}^{N}(t), \ \mathfrak{F}^{N}(t) = \sum_{j=1}^{I^{N}(0)} \lambda_{j}^{0}(t) + \sum_{i=1}^{A^{N}(t)} \lambda_{i}(t - \tau_{i}^{N}),$$

where $A_i^N(t) = 0$ if i is susceptible at time t, = 1 if i has been infected by time t.

• The jump time τ_i^N of $A_i^N(t)$ is the infection time of individual i, and

$$A_{i}^{N}(t) = \int_{0}^{t} \int_{0}^{\widetilde{\mathfrak{F}}^{N}(s^{-})} 1_{A_{i}^{N}(s^{-})=0} Q_{i}(ds, du),$$

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where the Q_i are i.i.d. PRMs.

- We want to show that $A^N(t) = \sum_{i=1}^{S^N(0)} A_i^N(t)$. Recall that $A^N(t) = \int_0^t \int_0^{S^N(s^-)} \widetilde{\mathfrak{F}}^{N(s^-)} Q(ds, du)$, while $A_i^N(t) = \int_0^t \int_0^{\widetilde{\mathfrak{F}}^{N(s^-)}} Q_i(ds, du) \wedge 1$.
- At the start of the epidemic, $\mathfrak{F}^N(s) = \sum_{j=1}^{I^N(0)} \lambda_j^0(s)$, which is independent of the Q_i 's. If we define Q(s,u) by

$$\begin{cases} Q_{1}(s,u), & \text{if } 0 \leq \frac{u}{\tilde{\mathfrak{F}}^{N}(s)} < 1; \\ Q_{2}(s,u-\bar{\mathfrak{F}}^{N}(s)), & \text{if } 1 \leq \frac{u}{\bar{\mathfrak{F}}^{N}(s)} < 2; \\ \dots, & \dots \\ Q_{S^{N}(0)}(s,u-(S^{N}(0)-1)\bar{\mathfrak{F}}^{N}(s)), & \text{if } S^{N}(0)-1 \leq \frac{u}{\bar{\mathfrak{F}}^{N}(s)} \leq S^{N}(0); \\ Q_{S^{N}(0)+1}(s,u), & \text{if } \frac{u}{\bar{\mathfrak{F}}^{N}(s)} > S^{N}(0); \end{cases}$$

we see that $A^N(t) = \sum_{i=1}^{S^N(0)} A_i^N(t)$ until the first jump, say $\tau_{(1)}^N$. That jump of $A^N(t)$ is a jump of A_i^N .

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• Starting from that jump, we suppress the i_1 line in the above definition of Q, and for all $j > i_1$, we replace the line

$$Q_j(s,u-(j-1)ar{\mathfrak{F}}^N(s)),$$
 if $j-1\leq rac{u}{ar{\mathfrak{F}}^N(s)} < j$ by $Q_j(s,u-(j-2)ar{\mathfrak{F}}^N(s)),$ if $j-2\leq rac{u}{ar{\mathfrak{F}}^N(s)} < j-1$.

• Now we have $A^N(t) = \sum_{1 \leq i \leq S^N(0), i \neq i_1} A_i^N(t)$, which is OK since the next jump is expected at a level $0 < u < (S^N(0) - 1)\overline{\mathfrak{F}}^N(s) = S^N(s)\overline{\mathfrak{F}}^N(s)$ until the next jump, etc.

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• Now suppose that we replace, for each $1 \le i \le S^N(0)$, $A_i^N(t)$ by $A_i(t)$ defined by

$$A_i(t) = \int_0^t \int_0^{\tilde{\mathfrak{F}}(s)} 1_{A_i(s^-)=0} Q_i(ds,du).$$

• We note that $\mathbb{P}(au_i>s)=\exp(-\int_0^sar{\mathfrak{F}}(r)dr)$. Recall that

$$\begin{split} \bar{\mathfrak{F}}(t) &= \bar{I}(0)\bar{\lambda}^0(t) + \int_0^t \bar{\lambda}(t-s)\bar{S}(s)\bar{\mathfrak{F}}(s)ds \\ &= \bar{I}(0)\bar{\lambda}^0(t) + \bar{S}(0)\int_0^t \bar{\lambda}(t-s)\exp(-\int_0^t \bar{\mathfrak{F}}(s)ds)\bar{\mathfrak{F}}(s)ds \\ &= \bar{I}(0)\bar{\lambda}^0(t) + \bar{S}(0)\mathbb{E}[\bar{\lambda}(t-\tau_i)] \,. \end{split}$$

• If we take the last line for the definition of $\bar{\mathfrak{F}}(t)$ in the formula for $A_i(t)$, we see that $A_i(t)$ solves a type of McKean-Vlasov SDE.

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• Now suppose that we replace, for each $1 \le i \le S^N(0)$, $A_i^N(t)$ by $A_i(t)$ defined by

$$A_i(t) = \int_0^t \int_0^{\widetilde{\mathfrak{F}}(s)} 1_{A_i(s^-)=0} Q_i(ds, du).$$

ullet We note that $\mathbb{P}(au_i>s)=\exp(-\int_0^sar{\mathfrak{F}}(r)dr).$ Recall that

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• If we take the last line for the definition of $\bar{\mathfrak{F}}(t)$ in the formula for $A_i(t)$, we see that $A_i(t)$ solves a type of McKean-Vlasov SDE.

• The point is that the sequence $\{A_i(\cdot), i \geq 1\}$ is i.i.d. So by the law of large numbers in D, as $N \to \infty$, a.s., locally unif. in t,

$$rac{1}{N}\sum_{i=1}^{S^N(0)}A_i(t)
ightarrow ar{S}(0)\int_0^t \exp(-\int_0^t ar{\mathfrak{F}}(s)ds)ar{\mathfrak{F}}(s)ds \ = \int_0^t ar{S}(s)ar{\mathfrak{F}}(s)ds$$

• It remains to prove that

Lemma

There exists a positive constant C_{T,λ^*} such that for all $N\geq 1,\, 0\leq t\leq T$

$$\frac{1}{N}\mathbb{E}\left[\sum_{i=1}^{S^N(0)}\sup_{0\leq t\leq T}|A_i^N(t)-A_i(t)|\right]\leq C_{T,\lambda^*}(\varepsilon_N+2N^{-1/2}),$$

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We have

$$\sup_{0\leq r\leq t}|A_i^N(r)-A_i(r)|\leq \int_0^t\int_{\tilde{\mathfrak{F}}(s)\wedge\tilde{\mathfrak{F}}^N(s^-)}^{\tilde{\mathfrak{F}}(s)\vee\tilde{\mathfrak{F}}^N(s^-)}Q_i(ds,du)$$

Hence

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$$\begin{split} & \mathbb{E}|\bar{\mathfrak{F}}(t) - \bar{\mathfrak{F}}^N(t)| \leq \mathbb{E}\frac{1}{N} \left| \sum_{j=1}^{I^N(0)} (\lambda_j^0(t) - \bar{\lambda}^0(t)) \right| \\ & + \frac{1}{N} \mathbb{E}\left| \sum_{i=1}^{S^N(0)} (\lambda_i(t - \tau_i^N) - \mathbb{E}[\lambda_i(t - \tau_i)]) \right| + \lambda^* \mathbb{E}[|\bar{I}^N(0) - \bar{I}(0)| + |\bar{S}^N(0) - \bar{S}(0)|] \end{split}$$

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Proof of the Lemma 2

We first note that, using second moment estimates,

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Time of extinction of an epidemic

- Suppose we are in a situation where, starting from $t=t_0$, an epidemic is declining (i.e. the mean number of susceptible individuals which an infected infects $R_{eff} < 1$), while the total number of infected individuals is M << N = the size of the population. Then deterministic models are no longer valid, the epidemic is well approximated by a sub-critical branching process, which decays essentially like an exponential $e^{-\rho(t-t_0)}$.
- We compare the time until extinction in 2 models : our VI infectivity model, and a Markovian SIR model with the same R_{eff} and ρ .
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Comparison with the time of extinction in a simplified model

• Comparison of the expectations of the times of extinction :

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Varying infectivity model	$\mathbb{E}[T_{ext}] \approx 18.7854$	$\mathbb{E}[T_{ext}] \approx 22.6568$
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